ECON42720 Causal Inference and Policy Evaluation 4b Matching and Re-weighting

Ben Elsner (UCD)

Approximate Matching

In most cases, we cannot find a perfect match for each treated unit

- Many variables are continuous
- We have many covariates
- ...and finite samples

Approximate matching allows us to match similar units

Approximate Matching: Questions to Ask Yourself



- 1: Which distance measure to use?
 - ► These determine how we measure similarity
- $2:\ \mbox{How to turn}$ distance into matches
 - Which matches are "good enough"?
 - Unique or multiple matches?
 - Cut-off points (calipers), number of neighbours?
- 3: How do we prune the data?
 - What do we do with units that are not matched?
- 4: Match with or without replacement?
 - Do we allow control units to be matched to multiple treated units?

Approximate Matching

There are two main methods for approximate matching:

- 1. **Distance Matching** \rightarrow minimise distance in X
- 2. Propensity Score Matching \rightarrow match on likelihood of being treated

A third type of matching is coarsened exact matching (CEM)

It is also possible to combine matching methods

Example: match exactly on some characteristics and approximately on others

Distance Matching: Questions to Ask

How do we measure distance, i.e. the similarity between treated and control units?

What is the **cut-off point** for a good match?

Do we consider multiple matches for each treated unit?

If so, what criterion determines which unit is a match?
And should each control unit get the same weight?

Starting point: treated units, with covariates age and age at which they left education



Treated and control units are different w.r.t. education



For each treated unit, we find the "closest" control unit in terms of X (Euclidean Distance)



Drop control units that are not close enough to any treated unit



Drop control units that are not close enough to any treated unit



Distance Matching: Nearest Neighbour Our estimation sample:



With one covariate, the distance is the Euclidean Distance

$$egin{aligned} |X_i - X_j|| &= \sqrt{(X_i - X_j)'(X_i - X_j)} \ &= \sqrt{\sum_{n=1}^k (X_{ni} - X_{nj})^2} \end{aligned}$$

For each treated unit, we find the control unit with the smallest distance $||X_i - X_i||$

Multiple Covariates: Mahalanobis Distance

With **multiple covariates** 1,..., k, we take into account the variance-covariance matrix $\hat{\Sigma}_X$ of the covariates

$$||X_i - X_j|| = \sqrt{(X_i - X_j)'\widehat{\Sigma}_X^{-1}(X_i - X_j)}$$

As before, for each treated unit, we find the control unit with the smallest distance $||X_i - X_j||$

Purpose of weighting with $\widehat{\Sigma}_X^{-1}$:

- Covariates become scale-invariant
- All distances are measured in terms of standard deviations

Nearest Neighbour Matching: Steps Involved

1. Preprocess (Matching)

Calculate the Mahalanobis Distance $||X_i - X_j|| = \sqrt{(X_i - X_j)'\widehat{\Sigma}_X^{-1}(X_i - X_j)}$

- Match each treated unit to the nearest control unit
- Prune control units if unused
- Prune matches if Distance>caliper (i.e. if they exceed a certain distance)
- 2. Estimation: calculate difference in means or run a regression

Other Distance Matching Methods

k-Nearest-neighbour Matching (NNM)

Match with the nearest neighbour or the k nearest neighbours in terms of X
 Take the average of these neighbours as the counterfactual

Radius and Kernel Matching

Match with all control units within a certain radius of the treated unit
 If all control units have equal weight, we call this radius matching
 If weights decay with distance, we call this kernel matching

Nearest Neighbour Matching with k = 3

Suppose you have a dataset with 2 treated and many control units



Nearest Neighbour Matching with k = 3

Now we select the three nearest neighbours for each treated unit; their average Y is the counterfactual for the treated unit



Nearest Neighbour Matching with k = 10

Now we select the ten nearest neighbours for each treated unit



Radius Matching

Here the researcher specifies a radius (r = 2) around the treated unit



Radius Matching

Each control unit within the radius has equal weight



Radius Matching: Larger Radius (r = 4)



Consider the following sample



Now suppose each control unit gets equal weight



Now let's use an Epanechnikov Kernel: further away \Rightarrow smaller weight



We want to create a weighted average by applying a kernel function

$$ar{Y} = rac{{\sum\limits_{i = 1}^n {{w_i}{Y_i}} }}{{\sum\limits_{i = 1}^n {{w_i}} }} = rac{{\sum\limits_{i = 1}^n {{\mathcal K}({X_i}){Y_i}} }}{{\sum\limits_{i = 1}^n {{\mathcal K}({X_i})} }}$$

There are many Kernel functions; they are typically concave and assign the highest weight to the smallest distance. Example: Epanechnikov kernel



K(X) is only defined between -1 and 1

Kernel Matching Within a Radius

We often use kernel weighting within a radius or the set of k nearest neighbours



More vs Better: the Bias-Variance Tradeoff

We often need to decide between unique matches and multiple control units

Researchers need to solve a bias-variance trade-off

Unique matches:

• Matches are precise but few \rightarrow low bias, high variance

Weighted average of multiple control units

Find many matches, but these are imprecise \rightarrow high bias, low variance

Idea of CEM:

- Coarsen X (for example different age groups)
- Perform exact matching based on coarsened data

Advantage: easy and fast

Disadvantages:

researcher degrees of freedom (categories are chosen by the researcher)
 curse of dimensionality (few categories: many but imprecise matches; many categories: few but more precise matches)

Starting point: same as before



Coarsen: divide variables into categories



Now see which cells contain treated and control units



We find matches within cells



We find matches within cells

We need to take a stand regarding matching within the cells

- nearest neighbour or k nearest neighbours
- with or without replacement
- kernel distance function (usually not necessary)

Idea: predict the probability that a given unit is treated based on X

• The probability $Pr(D_i = 1|X)$ is called the **propensity score**

Match units with a similar probability of being treated (propensity score)

Estimate the ATT based on the matched dataset

Starting point: treated and control units



For each unit, we want to predict the probability that it is treated based on X



Let's do this: predict the probability of being treated



For each treated unit, find the control unit with the closest propensity score



Prune control units that have not been matched



The result is your matched dataset



The result is your matched dataset



We can now regress the outcome on the treatment status



Estimation of the propensity score

The propensity score must satisfy the **balancing property**. It implies:

- Observations with the same propensity score have the same distribution of observable covariates independently of treatment status;
- ► For a given propensity score **assignment to treatment is random**, hence treated and control units are on average observationally identical.

Estimation of the propensity score

Can use any **standard probability model** to estimate the propensity score, e.g. a logit model:

$$Pr\{D_i=1|X_i\}=rac{e^{\lambda h(X_i)}}{1+e^{\lambda h(X_i)}},$$

where $h(X_i)$ is a function of covariates with linear and higher order terms.

Inverse Probability Weighting (IPW)

IPW is based on the propensity score. Ingredients:

- The propensity score of being treated: p(X)
- The propensity score of beign untreated: 1 p(X)

Units are weighted by the inverse propensity score of THEIR treatment status

Why does this work?

- Weights "create" similar observations in terms of X
- Treated observations with similar X as untreated get a high weight because they are similar

Inverse Probability Weighting and the ATE

It can be shown that IPW identifies the ATE in the population:

$$\Delta = E \left[\mu_1(X) - \mu_0(X) \right] = E \left[\frac{E[Y \cdot D \mid X] \cdot D}{p(X)} - \frac{E[Y \cdot (1 - D) \mid X] \cdot (1 - D)}{1 - p(X)} \right]$$
$$= E \left[\frac{Y \cdot D}{p(X)} - \frac{Y \cdot (1 - D)}{1 - p(X)} \right]$$

Inverse Probability Weighting and the ATT

The ATT is identified by

$$\Delta_{D=1} = E\left[\frac{Y \cdot D}{\Pr(D=1)} - \frac{Y \cdot (1-D) \cdot p(X)}{(1-p(X)) \cdot \Pr(D=1)}\right]$$

IPW Weights Example

Observation	Treated	PS(p(x))	weight
1	1	0.6	$\frac{1}{p(x)} = \frac{1}{0.6} = 1.67$
2	0	0.6	$\frac{1}{1-p(x)} = \frac{1}{0.4} = 2.5$
3	1	0.9	$\frac{1}{p(x)} = \frac{1}{0.9} = 1.11$
4	0	0.9	$\frac{1}{1-\rho(x)} = \frac{1}{0.1} = 10$

Here, observations 2 and 4 (untreated) get fairly large weights because they have similar X to typical treated units

Observation 3 (treated) gets a low weight because it is dissimilar to most untreated units

Matching and Causal Inference

Matching is NOT a causal identification strategy

- Neither is regression
- It is a data reduction/pre-processing technique

It helps us to achieve **balance on observables** X

Causal identification rests on the conditional independence assumption

- ▶ given X, D should be as good as randomly assigned
- ▶ i.e. X has to capture all confounders

Matching and Causal Inference: a DAG

If this is the correct DAG, our matching needs to account for A, B and C



PSM and Causal Inference

With PSM, the additional assumption is that the **propensity score is correctly specified and closes all backdoor paths**



Matching Example: Broockman (2013)

Research question: are black politicians more likely to help black citizens even if the incentives are low?

Methodology: audit study; sent emails to U.S. state legislators; asking them to help them sign up for unemployment benefits

Experimental variation:

Sender with black vs. white name

Sender lives in same district as legislator or far away

Matching: white and black legislators with similar characteristics

Step 1: Check for Common Support



Left: share of black voters in the district

Right: distribution of propensity scores

propensity score p (probability of being black) based on share of black voters, median household income, legislator is a democrat

Step 1: Common Support

Problem: areas with high p have no white legislators, areas with low p have no black legislators

Solution: prune areas without common support (at least 10 control obs within a .02 bin)



Balance in Broockman (2013)

Broockman performs Mahalanobis (nearest neighbour) matching; here is the balancing table

Before Matching	After Matching
3.33	3.333
4.435	3.316
-97.057	1.455
<.0001	0.164
0.517	0.515
0.063	0.513
224.74	1.288
<.0001	0.034
0.978	0.978
0.501	0.978
325.14	0
<.0001	1
	Before Matching 3.33 4.435 -97.057 <.0001 0.517 0.063 224.74 <.0001 0.978 0.501 325.14 <.0001

Careful when Performing Balancing Tests

Focusing just on mean differences can be deceptive. Consider these two distributions:



Always check the full distribution of covariates before and after matching

Full Distribution of X and p after IPW



Not perfect, but not so bad either...

Data preparation in R

We use the excellent Matching package in R. A great alternative is MatchIt

```
library(Matching)
library(causaldata)
library(tidyverse)
```

br <- causaldata::black_politicians</pre>

```
# Outcome
Y <- br %>%
  pull(responded)
# Treatment
D <- br %>%
  pull(leg black)
# Matching variables
# Note select() is also in the Matching package, so we specify dplyr
X <- br %>%
  dplyr::select(medianhhincom, blackpercent, leg democrat) %>%
  as.matrix()
```

Mahalanobis distance matching in R

```
# Set weight=2 for Mahalanobis distance
M <- Match(Y, D, X, Weight = 2, caliper = 1)</pre>
```

See treatment effect estimate
summary(M)

##

##	Estimate	-0.0073462	
##	AI SE	0.072683	
##	T-stat	-0.10107	
##	p.val	0.91949	
##			
##	Original numb	er of observations	5593
##	Original numb	per of treated obs	364
##	Matched number	er of observations	363
##	Matched numbe	er of observations (unweighted).	405
##			
##	Caliper (SDs)		111
##	Number of obs	dropped by 'exact' or 'caliper'	1

Previous slide: the estimate -0.007346 means that black legislators were 0.7 percentage points less likely to respond to emails

This effect is not statistically significant

$Comparison \ with \ OLS$

Table 2

	Dependent variable: responded		
	(1)	(2)	
leg_black	-0.032	-0.035	
	(0.027)	(0.039)	
medianhhincom		0.014***	
		(0.005)	
blackpercent		0.081	
		(0.063)	
leg democrat		-0.039***	
		(0.014)	
Constant	0.425***	0.377***	
	(0.007)	(0.025)	
Observations	5,593	5,593	

The covariate plots showed that there is little common support

Matching rests on comparable observations with common support in X

 $\ensuremath{\mathsf{OLS}}$ uses observations without common support; this explains the difference in the estimates

Matching and Re-Weighting: Conclusion

Matching and re-weighting are data subsetting techniques

- They help to achieve balance on observables X
- They are particularly useful when there is little common support

Matching and re-weighting are not causal identification strategies

- we need to rely on the conditional independence assumption to identify causal effects
- whether this holds depends on the context
- even the best matching procedure in the world cannot fix a bad research design

Matching and Re-Weighting: Conclusion

There are many matching procedures

- It is easy to get lost in the details
- ▶ The most important thing is to **achieve balance on observables** *X*
- When choosing a method, keep the bias-variance tradeoff in mind

Showing robustness to different matching procedures is very important

Matching in R

R has many packages for matching; most of them do similar things but have their strengths and weaknesses. Here are some very good ones:

- Matching
- MatchIt
- cem for Coarsened Exact Matching
- optmatch

MatchIt covers all the bases

References

Broockman, David E. 2013. Black Politicians Are More Intrinsically Motivated to Advance Blacks' Interests: A Field Experiment Manipulating Political Incentives. American Journal of Political Science, 57(3), 521–536.

APPENDIX



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benjamin.elsner@ucd.ie



www.benjaminelsner.com



Sign up for office hours



YouTube Channel



@ben_elsner

LinkedIn

Contact

Prof. Benjamin Elsner

University College Dublin School of Economics Newman Building, Office G206 benjamin.elsner@ucd.ie

Office hours: book on Calendly